

Childhood Cancer Survivor Study Analysis Concept Proposal

August 9, 2011

1. **Title:** Development and validation of an absolute risk prediction model for thyroid cancer in childhood cancer survivors

2. **Investigators:** This proposed study will be within the Second Malignancies Working Group. Proposed investigators include:

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3. **Background and Rationale**

Cancer of the thyroid gland is relatively rare, although the incidence has significantly increased in the past two decades, in particular for papillary tumors (Enewold et al, 2009). In contrast to most other solid cancers, thyroid cancer is more common in females than males and the average age at diagnosis is in the forties rather than later in life. A well-established thyroid cancer risk factor is exposure to ionizing radiation in the head and neck region at a young age. Others are iodine deficiency and a history of thyroid diseases, particularly nodular hyperplasia and thyroid adenomas (reviewed in Ron 1996).

Several studies of childhood cancer survivors that included children whose thyroid gland was exposed to high radiation doses (>10 Gy) (Tucker 1991; deVathaire 1999, Sigurdson 2005; Ronckers 2006) have found increased risk of thyroid cancer. The relationship with radiation dose has been confirmed as curvilinear (Ronckers 2006; Bhatti 2010) using data from the Childhood Cancer Survivor Study (CCSS) which is the largest cohort study to report on thyroid cancer risk. Ronckers et al (2006) used a case-control study design with 69 thyroid cancer cases and found a suggestion that age at exposure modified the relation with dose. Bhatti et al (2010) used a cohort design and based on 119 cases confirmed the curvilinear dose-response relationship. Bhatti et al (2010) further found that age at exposure modified the excess relative risk and that gender and time since exposure modified the excess absolute risk of subsequent thyroid cancer. Most recently we have found that alkylating agents also increased thyroid cancer risk, but only in the radiation dose range under 20 Gy, with a suggestive effect for an individual alkylating agent, procarbazine (Veiga submitted). Analyses are currently underway to combine all the childhood cohorts from the US and Europe in a pooled study to further elucidate the shape of the radiation dose-response and risk modifiers (Pooled International Radiation And Thyroid cancer Epidemiology Study, PIRATES-led by Drs. Veiga and Lubin) (Veiga, in preparation).

Given the high risk associated with radiation treatment in individuals exposed at young ages and the curvilinear relationship with dose (unusual for most second solid tumors), predicting absolute risk of thyroid cancer in childhood cancer survivors is complex but would be useful for physicians when they are making screening recommendations and counseling patients. Our goal is thus to build a model to predict absolute thyroid cancer risk among childhood cancer patients who have survived to adulthood when the risk of thyroid cancer continues to be elevated, as elevated thyroid cancer risk from radiation treatment is present for decades after first exposure (reviewed in Ron 1996). To our knowledge, no such model has been developed to date.

4. Specific aims/research objectives/hypotheses to be tested

While including detailed information about a patient's treatment history including absorbed thyroid dose is expected to give the most accurate risk prediction, such a model might have limited practical use if treatment information is not readily available to the clinician. We thus propose to

build two models to predict risk of any second primary thyroid cancer in individuals who were diagnosed with a childhood cancer and survived at least 5 years past that diagnosis.

1. Model 1 will include variables that were obtained through self report in the studies, such as age at first cancer diagnosis, gender, type of first primary cancer, and thyroid conditions. A variable selection procedure will be applied to identify important risk factors among the self-reported variables which will be based on multivariable analyses with possible first order interactions.
2. Model 2 will consider the additional inclusion of more detailed information about treatment history, including yes/no indicators for the receipt of radiation or chemotherapy treatment, type of chemotherapy received, and body regions irradiated. The same variable selection procedure will be applied to identify important risk factors in multivariate analyses of the treatment and self-report variables. We regard Model 2 as a “clinical” model in contrast to the self-report model of Model 1 since treatment history for the first cancer might require chart review.
3. The “gold standard model” will include dose estimates to the thyroid (from the UTMDACC radiation dosimetry group) for radiation treatment of the first cancer, irradiated body region and type of chemotherapy in addition to information on self reported risk factors. Again, variable selection will be applied to determine important variables.
4. We will validate all models in an independent cohort and compare their predictive performance.

5. Analysis Framework:

Model building

The absolute risk A^* of developing thyroid cancer (TC) in the age interval (a,b) is given by

$$A^*(a,b) = \int \lambda_t(u,x) S^*(u-) du / S^*(a-) \quad (1)$$

where $S^*(a) = \exp(-\int \{ \lambda_t(u, x) + \lambda_o(u, x) + \lambda_M(u) \} du)$.

This models the incidence of secondary primary TC, in the presence of competing risks from all other secondary primary cancers and causes of death other than cancer captured by the hazards $\lambda_o(u, x)$ and $\lambda_M(u)$ respectively. Let T denote the age of onset of the TC outcome, age at a second cancer diagnosis other than TC or age at death due to other causes. The cause specific hazards that may depend on covariates x are defined as $\lambda_t(a, x) = \lim_{\epsilon \rightarrow 0} P(a \leq T < a + \epsilon, J=TC, x)/\epsilon$. We model $\lambda_t(a, x) = \lambda_o(a) rr(a, x)$ as the product of the age specific hazard rate and the relative risk part for TC, $rr(a, x)$, that includes covariates. We may also include covariates in the hazard estimation for the competing events of other second primary cancer and death. For the patient based and clinical models the relative risks will be based on an exponential model, i.e. $rr(a, x) = \exp(\beta X)$, with possible interactions of X with age, and we will use an excess relative risk-based model that includes known (estimated from UTMDACC) dose for the gold standard model. We propose to combine data from the Childhood Cancer Survivor Study (CCSS) (Bhatti et al., 2010), the Nordic study (Svahn-Tapper et al., 2006), and the Late Effects Study Group (LESG) (Tucker et al., 1991) to estimate relative risk parameters for the models to minimize variability of the estimates, after assessing heterogeneity of the estimates across the three studies. We will do this by combining the cohort partial likelihood for CCSS with a logistic regression likelihood for the case-control studies and jointly maximizing the combined likelihood. Table 1 lists the number of cases and non-cases available for each study. We will obtain semiparametric estimates of the baseline hazard rates for thyroid cancer and competing events based on the event times within the CCSS cohort. We will extend variance computations based on influence function methods (Graubard and Fears, 2005; Pfeiffer and Petracchi, 2011) to the cohort setting and when relative risk parameters are estimated from cohort and case-control data combined. Variables for the models will be selected based on statistical significance and also biological relevance.

Model validation

We propose to validate the models using the French-UK CCSS (de Vathaire et al., 1999) and to compare the performance of all models in the validation data. We will compare the expected (E) and the observed (O) numbers of cases overall and in subgroups defined by risk-factor

combinations. The expected number of cases is calculated by summing the individual projected probabilities, given the baseline covariate values for each person over the time from entry into the validation cohort to end of follow-up. The 95% confidence intervals for E/O ratios are calculated

using the normal approximation to the Poisson distributions: $\frac{E}{O} e^{\pm 1.96\sqrt{\frac{1}{O}}}$. An E/O ratio above 1

indicates the model overestimates the risk of cancer and $E/O < 1$ indicates the model underestimates the risk of cancer. We will evaluate the discriminatory accuracy of the prediction models using the area under the receiver-operating characteristic curve (AUC).

After comparing the predictive performance of the models we will decide if a risk prediction tool based on Model 1 or Model 2 might be clinically useful.

Proposed tables follow:

Table 1: Description of analytic data set, cases and non-cases, and risk factor distribution for CCSS, LESG, and Nordic data.

| Characteristic | Cases, No. (%) | | | Non-cases, No. (%) | | |
|---|----------------|-----------|-----------|--------------------|-----------|-----------|
| | CCSS | LESG | NORDIC | CCSS | LESG | NORDIC |
| N | 113 | 22 | 13 | 11868 | 82 | 36 |
| Gender: | | | | | | |
| Male | 34 (30.1) | 8 (36.4) | 3 (23.1) | 6265 (52.8) | 28 (34.1) | 9 (25.0) |
| Female | 79 (69.9) | 14 (63.6) | 10 (76.9) | 5603 (47.2) | 54 (65.9) | 27 (75.0) |
| First Primary Diagnosis | | | | | | |
| Leukemia | 27 (23.9) | NA | 2 (15.4) | 4061 (34.2) | NA | 7 (19.4) |
| CNS | 13 (11.5) | 1 (4.5) | 0 (0.0) | 1543 (13.0) | 3 (3.7) | 7 (19.4) |
| HD | 39 (34.5) | 5 (22.7) | 5 (38.5) | 1503 (12.7) | 15 (18.3) | 1 (2.8) |
| NHL | 6 (5.3) | 2 (9.1) | 1 (7.7) | 876 (7.4) | 5 (6.1) | 1 (2.8) |
| Kidney (Wilms) | 2 (1.8) | 4 (18.2) | 1 (7.7) | 1045 (8.8) | 21 (25.6) | 2 (5.6) |
| Neuroblastoma | 8 (7.1) | 7 (31.8) | 0 (0.0) | 805 (6.8) | 27 (32.9) | 1 (2.8) |
| Soft tissue sarcoma) | 8 (7.1) | 1 (4.5) | 0 (0.0) | 1033 (8.7) | 3 (3.7) | 2 (5.6) |
| Bone | 10 (8.8) | 0 (0.0) | 1 (7.7) | 1002 (8.4) | 4 (4.9) | 3 (8.3) |
| Age at diagnosis of first primary (years) | | | | | | |
| <5 | 28 (24.8) | 11 (50.0) | 3 (23.1) | 4848 (40.8) | 51 (62.2) | 12 (33.3) |
| 5-9 | 24 (21.2) | 5 (22.7) | 2 (15.4) | 2602 (21.9) | 15 (18.3) | 6 (16.7) |
| 10-14 | 48 (42.4) | 4 (18.2) | 2 (15.4) | 2339 (19.7) | 11 (13.4) | 4 (11.1) |
| 15+ | 13 (11.5) | 2 (9.1) | 6 (46.2) | 2079 (17.5) | 5 (6.1) | 14 (38.9) |
| Follow-up (years) | | | | | | |
| 5-14 | | | | | | |
| 15-24 | | | | | | |

| | | | | | | |
|--|--|--|--|--|--|--|
| 25+ | | | | | | |
| Age at last follow-up | | | | | | |
| Cutpoints TBD | | | | | | |
| Treatment Era or Birth year | | | | | | |
| Cutpoints TBD | | | | | | |
| Visits to a physician | | | | | | |
| None | | | | | | |
| Number of times, other cutpoints TBD | | | | | | |
| Number of years since last physical examination | | | | | | |
| Smoking | | | | | | |
| Former | | | | | | |
| Current | | | | | | |
| Never | | | | | | |
| Use of thyroid hormone replacement | | | | | | |
| No | | | | | | |
| Yes | | | | | | |
| History of underactive Thyroid | | | | | | |
| No | | | | | | |
| Yes | | | | | | |
| Not sure | | | | | | |
| History of overactive Thyroid | | | | | | |
| No | | | | | | |
| Yes | | | | | | |
| Not sure | | | | | | |
| History of thyroid gland enlargement | | | | | | |
| No | | | | | | |
| Yes | | | | | | |
| Not sure | | | | | | |
| Chemotherapy | | | | | | |
| No | | | | | | |
| Yes | | | | | | |
| Chemotherapy | | | | | | |
| No | | | | | | |
| Yes | | | | | | |
| Chemotherapy drugs | | | | | | |
| Alkylating agents | | | | | | |
| Anthracyclines | | | | | | |
| Bleomycin | | | | | | |
| Platinum | | | | | | |
| Radiation | | | | | | |
| No | | | | | | |
| Yes | | | | | | |
| Radiation Dose (Gy) (categories subject to change) | | | | | | |
| None | | | | | | |
| >0-<5 | | | | | | |

| | | | | | | |
|--------|--|--|--|--|--|--|
| 5-10 | | | | | | |
| >10-20 | | | | | | |
| >20-30 | | | | | | |
| >30-40 | | | | | | |
| >40 | | | | | | |

TBD=To be determined

Table 2a: Multivariate relative risk estimates for patient-based model (factors TBD)

Variables that will be assessed during model building include: Age at first primary, birth year, gender, time since first cancer, type of first cancer, thyroid nodule, thyroid enlargement, under- or over-active thyroid, taking thyroid medication, and personal screening practices and doctor visits (may increase thyroid cancer detection).

Table 2b: Multivariate relative risk estimates for patient- based and clinician model (may include a qualitative expression of dose if such is available for the tumor and tumor location—see appendix for possible combinations)

Table 3: Examples of absolute risk estimates for select risk factor profiles

Table 4: Observed and expected thyroid cancers in cells of covariates in the validation cohort

6. Special Considerations

In 2006, in our concept proposal that was approved by the CCSS publications committee, we had described building a prediction model nicknamed by Dr. Anna Meadows as the “yardstick”, with the intent that it could be used in clinical practice during survivor follow-up visits. In the intervening years cancer risk prediction modeling has gained in sophistication and accuracy, partly due to the efforts of Dr. Ruth Pfeiffer (Freedman 2009; Park 2009; Pfeiffer 2010; O’Brien 2011). We have been fortunate to engage Dr. Pfeiffer to oversee the risk prediction effort for thyroid cancer along with her post-doctoral fellow, Dr. Stephanie Kovalchik.

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Appendix from 2006 proposal:

Possible selection of representative/typical patients to illustrate effects of various patient and treatment characteristics on estimated thyroid dose

| Patients with XRT | | | Age at Diagnosis | | | | Thyroid Radiation Exposure |
|----------------------------------|--------------------|--------------------------|----------------------------------|----------|----------|----------|----------------------------|
| | | | Number of patients (% age group) | | | | |
| CCSS Primary Disease | % of Primary Group | Total number of patients | <5 | 5 to 9 | 10 to 14 | 15 to 20 | |
| Leukemia | | 2913 | | | | | |
| Spine XRT | 15% | 439 | 243(16%) | 113(15%) | 66(15%) | 17(8%) | high |
| No Spine XRT | 84% | 2474 | 1241(84%) | 655(85%) | 379(85%) | 199(92%) | low |
| CNS | | 1153 | | | | | |
| C-Spine XRT | 39% | 453 | 147(39%) | 159(45%) | 107(37%) | 40(31%) | high |
| No C-Spine XRT | 61% | 700 | 234(61%) | 194(55%) | 183(63%) | 89(69%) | medium |
| HD | | 1567 | | | | | |
| Above the diaphragm | 95% | 1496 | 36(84%) | 170(89%) | 512(96%) | 778(97%) | medium-high |
| No treatment above the diaphragm | 5% | 71 | 7(16%) | 21(11%) | 20(4%) | 23(3%) | low |
| NHL | | 636 | | | | | |
| Above the diaphragm | 57% | 365 | 39(42%) | 116(58%) | 115(58%) | 95(65%) | medium-high |
| No treatment above the diaphragm | 43% | 271 | 53(58%) | 83(42%) | 83(42%) | 52(35%) | low |
| Wilms/Kidney | | 693 | | | | | |
| Chest treated | 38% | 263 | 157(33%) | 93(51%) | 11(41%) | 2(29%) | medium |
| No chest treatment | 62% | 430 | 318(67%) | 91(50%) | 16(59%) | 5(71%) | low |
| Neuroblastoma | | 412 | | | | | |
| Chest or neck | 47% | 195 | 165(46%) | 21(57%) | 8(73%) | 1(25%) | medium-high |
| No chest or neck | 53% | 217 | 195(54%) | 16(43%) | 3(27%) | 3(75%) | low |
| Soft Tissue Sarcoma | | 686 | | | | | |
| Chest or neck or face | 58% | 396 | 141(57%) | 119(73%) | 76(53%) | 60(46%) | medium-high |
| No chest or neck or face | 52% | 290 | 105(43%) | 45(27%) | 68(47%) | 72(55%) | low |
| Bone tumors | | 382 | | | | | |
| Chest or neck or face | 48% | 197 | 13(65%) | 38(49%) | 73(53%) | 61(42%) | medium-high |
| No chest or neck or face | 52% | 185 | 7(35%) | 40(51%) | 65(47%) | 85(58%) | low |
| TOTAL | | 8442 | | | | | |